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# RIFFLE SHUFFLES WITH BIASED CUTS

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Stanford University



August 1, 2012

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**Theorem.** (Bayer–Diaconis) On  $\mathfrak{S}_n$ , we have  $P_{\frac{1}{a}} * P_{\frac{1}{b}} = P_{\frac{1}{ab}}$ .

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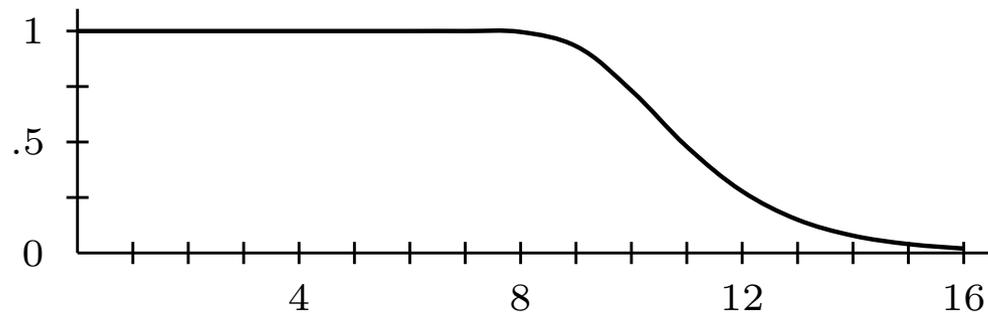
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$k$	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16
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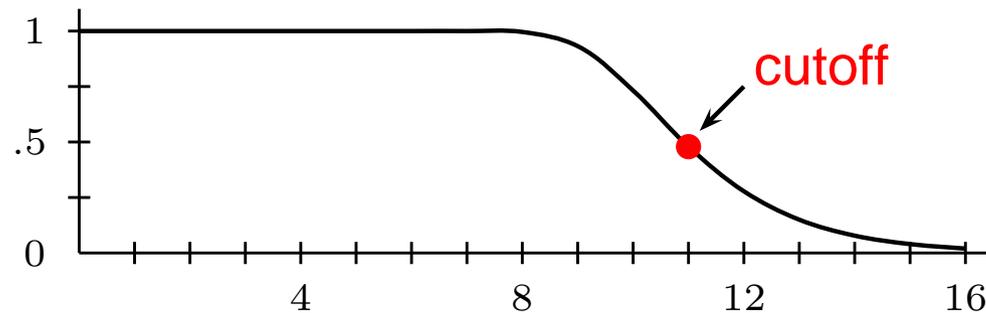
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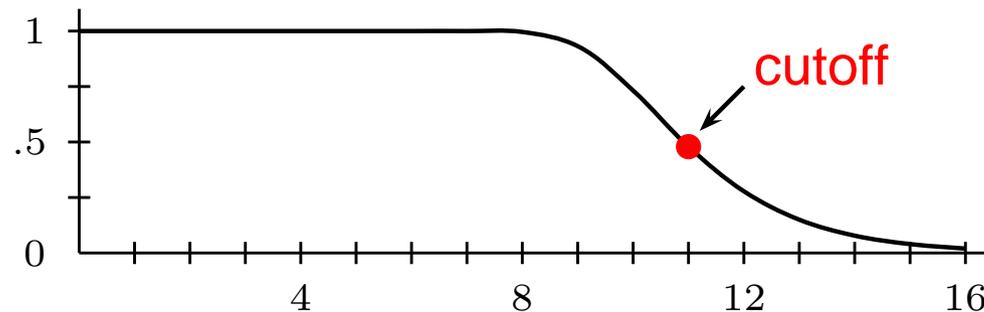
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Precisely, for  $a_n, b_n \rightarrow \infty$  with  $b_n/a_n \rightarrow 0$ , the chains  $P_n, \pi_n$  satisfy an  $a_n, b_n$  cutoff if for all real fixed  $\theta$  with  $k_n = \lfloor a_n + \theta b_n \rfloor$

$$\|P_n^{k_n} - \pi_n\| \longrightarrow c(\theta) \quad \text{where} \quad \begin{cases} c(\theta) \rightarrow 0 & \text{as } \theta \rightarrow \infty \\ c(\theta) \rightarrow 1 & \text{as } \theta \rightarrow -\infty \end{cases}$$

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Repeated  $\theta$ -shuffles convolve:  $\theta = (\theta_1, \dots, \theta_a)$  and  $\eta = (\eta_1, \dots, \eta_b)$ ,

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The **monomial** quasisymmetric function basis is (**compositions**)

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Gessel's **fundamental** quasisymmetric function basis is (**subsets**)

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The fundamental basis is related to the monomial basis by

$$Q_{D(\beta)}(X) = \sum_{\alpha \text{ refines } \beta} M_\alpha(X)$$

# ***Biased distribution***

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**Theorem.** (Fulman, Stanley) For  $\sigma \in \mathfrak{S}_n$  and  $\theta = (\theta_1, \theta_2, \dots, \theta_a)$ ,

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**Corollary.** On  $\mathfrak{S}_n$ , we have  $\text{SEP}(P_{\theta}) = 1 - n!Q_{[n-1]}(\theta)$ .

# ***Biased separation***

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**Theorem.** (A-D-S) On  $\mathfrak{S}_n$ , the separation distance for  $P_\theta$  is

$$\text{SEP}(k) = 1 - n! \left( \sum_{\lambda \vdash n} (-1)^{n-\ell(\lambda)} z_\lambda^{-1} \prod_{i=1}^n p_i(\boldsymbol{\theta})^{kn_i(\lambda)} \right)$$

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**Proof:**  $e_n(X) = Q_{[n-1]}(X)$ ,

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A similarly sharp analysis for **total variation** remains open for  $\theta \neq \frac{1}{2}$ .

## References

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- J. Fulman. The combinatorics of biased riffle shuffles. *Combinatorica*, 18(2):173–184, 1998.
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